

Random walk

Specific model for a walker in 1-d

- Step length l
- step interval T
- probability of moving to right p
- probability of moving to left q
- $p+q=1$

Let there be N_1 steps to right & N_2 steps to the left, s.t.,

$$N = N_1 + N_2 \equiv \text{total number of steps.}$$

and, $m = N_1 - N_2 \equiv \text{displacement to the right (in units of } l\text{).}$

Therefore, the probability that the walker has a displacement mL after N steps, $P_N(m) = \frac{N!}{N_1! N_2!} p^{N_1} q^{N_2} = \frac{N!}{\underbrace{(N+m)!}_{W_N(N_1)} \underbrace{(N-m)!}_{(N-m)!}} p^{\frac{N+m}{2}} q^{\frac{N-m}{2}}$

$$\langle N_1 \rangle = \sum_{N_1=0}^N N_1 \frac{N!}{N_1! N_2!} p^{N_1} q^{N_2} = p \frac{\partial}{\partial p} \left\{ \sum_{N_1=0}^N \frac{N!}{N_1! N_2!} p^{N_1} q^{N_2} \right\}$$

$$\text{But, } \sum_{N_1=0}^N \frac{N!}{N_1! N_2!} p^{N_1} q^{N_2} = (p+q)^N.$$

$$\Rightarrow \langle N_1 \rangle = Np(p+q)^{N-1} = Np \quad (\because p+q=1.)$$

$$\text{Similarly, } \langle N_1^2 \rangle = \left(p \frac{\partial}{\partial p} \right)^2 (p+q)^N$$

$$= Np(p+q)^{N-1} + N(N-1)p^2(p+q)^{N-2}$$

$$= Np + N(N-1)p^2$$

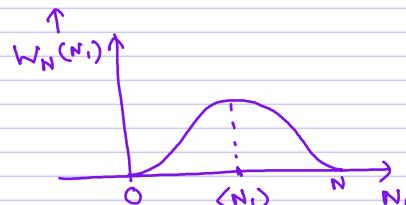
$$\Rightarrow \langle (\Delta N_1)^2 \rangle = \langle N_1^2 \rangle - \langle N_1 \rangle^2 = Np + N(N-1)p^2 - N^2p^2$$

$$= Np(1-p) = Npq.$$

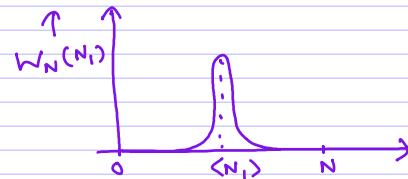
$$\Rightarrow \sqrt{\langle (\Delta N_1)^2 \rangle} = \sigma_{N_1} = \sqrt{Npq}.$$

$$\therefore \frac{\sigma_{N_1}}{\langle N_1 \rangle} = \frac{\sqrt{Npq}}{Np} = \frac{1}{\sqrt{N}} \sqrt{\left(\frac{q}{p}\right)}.$$

$$\text{Thus, for } q=p=\frac{1}{2} \rightarrow \frac{\sigma_{N_1}}{\langle N_1 \rangle} = \frac{1}{\sqrt{N}}.$$



$$\sigma_{N_1} \sim \langle N_1 \rangle$$



$$\sigma_{N_1} \ll \langle N_1 \rangle$$

Generating function for Binomial distribution

$$f(x) = \sum_{N_1=0}^N W_N(N_1) x^{N_1} = (px+q)^N.$$

Partially differentiating w.r.t. x on both sides, and

$$\Rightarrow \left. \sum_{N_1=0}^N N_1 x^{N_1-1} W_N(N_1) \right|_{x=1} = \left. N p (px+q)^{N-1} \right|_{x=1}.$$

$$\Rightarrow \langle N_1 \rangle = Np.$$

Similarly, taking a second partial derivative w.r.t. x and setting $x=1$,

$$\Rightarrow \langle N_1(N_1-1) \rangle = N(N-1)p^2.$$

$$\therefore \langle N_1^2 \rangle - \langle N_1 \rangle = N(N-1)p^2$$

$$\Rightarrow \langle N_1^2 \rangle = N^2 p^2 - Np^2 + Np \quad (\because \langle N_1 \rangle = Np).$$

$$\therefore \sigma_{N_1}^2 = (\Delta N_1)^2 = \langle N_1^2 \rangle - \langle N_1 \rangle^2 = N^2 p^2 - Np^2 + Np - N^2 p^2 \\ = Np(1-p) = Npq.$$

$$\Rightarrow \sigma_{N_1} = \sqrt{Npq}.$$

Binomial distribution \rightarrow Gaussian distribution

Recap: Stirling approximation

$$n! \sim n^n e^{-n} \sqrt{2\pi n} \quad (\text{large } n)$$

$$n! = \int_0^\infty x^n e^{-x} dx = \int_0^\infty e^{n(\ln x - x)} dx$$

$$\text{Let } f(x) = n(\ln x - x).$$

$$f'(x) = \frac{n}{x} - 1.$$

$$f''(x) = -\frac{n}{x^2}.$$

$$f(n) = n \ln n - n$$

$$f'(n) = 0$$

$$f''(n) = -\frac{1}{n}$$

Expanding $f(x)$ in a Taylor series about n ,

$$\Rightarrow f(x) \approx n \ln n - n - \frac{1}{2n}(x-n)^2.$$

$$\therefore n! \sim \int_0^\infty e^{n \ln n} e^{-n} e^{-(x-n)^2/2n} dx = n^n e^{-n} \int_0^\infty e^{-\frac{(x-n)^2}{2n}} dx$$

$$\text{Let } x-n=y \Rightarrow dx = dy$$

$$\text{For } x=0, y=-n$$

$$\text{For } x=\infty, y=\infty$$

$$\therefore \text{for large enough } n$$

$$n! \sim n^n e^{-n} \int_{-n}^{\infty} e^{-y^2/2n} dy$$

$$\sim \sqrt{2\pi n} n^n e^{-n}.$$

$$\therefore \ln n! \approx n \ln n - n.$$

$$W_N(N_1) = \frac{N!}{N_1!(N-N_1)!} p^{N_1} q^{N-N_1}$$

$$\Rightarrow \ln W_N(N_1) = \ln N! - \ln N_1! - \ln(N-N_1)! + N_1(\ln p + (N-N_1)\ln q).$$

$$= N \ln N - N_1 \ln N_1 - (N-N_1) \ln(N-N_1) + N_1(\ln p + (N-N_1)\ln q).$$

$$\therefore \left. \frac{\partial \ln W_N(N_1)}{\partial N_1} \right|_{N_1=\tilde{N}_1} = -1 - \ln \tilde{N}_1 + 1 + \ln(N-\tilde{N}_1) + \ln p - \ln q.$$

$$= 0 \Rightarrow \frac{N-\tilde{N}_1}{\tilde{N}_1} = \frac{q}{p} \Rightarrow \tilde{N}_1 = Np.$$

i.e., most probable value equals average value.

$$(\tilde{N}_1 = \langle N_1 \rangle = Np)$$

$$\text{Similarly, } \frac{\partial^2 \ln W_N(N_1)}{\partial N_1^2} = -\frac{1}{N_1} - \frac{1}{N-N_1}$$

$$\therefore \left. \frac{\partial^2 \ln W_N(N_1)}{\partial N_1^2} \right|_{N_1=\tilde{N}_1} = -\frac{1}{Np} - \frac{1}{Nq} = -\frac{1}{Npq}.$$

∴ Using Taylor series expansion,

$$\ln W_N(N_1) \approx \ln W_N(\tilde{N}_1) - \frac{1}{2Npq}(N_1 - \tilde{N}_1)^2$$

$$\therefore W_N(N_1) \approx W_N(\tilde{N}_1) \exp \left\{ -\frac{(N_1 - \langle N_1 \rangle)^2}{2\sigma_{N_1}^2} \right\}$$

(Gaussian distribution).

Alternate derivation:

$$W_N(N_1) = \frac{N!}{N_1!(N-N_1)!} p^{N_1} q^{N-N_1}.$$

Use Stirling approximation, $n! \sim n^n e^{-n} \sqrt{2\pi n} (1 + O(\frac{1}{n}))$

$$\begin{aligned} \Rightarrow W_N(N_1) &\approx \frac{N^N e^{-N} \sqrt{2\pi N}}{N_1^{N_1} e^{-N_1} \sqrt{2\pi N_1} (N-N_1)^{N-N_1} e^{-(N-N_1)} \sqrt{2\pi (N-N_1)}} (1 + O(\frac{1}{N})) \\ &= \frac{N^N}{N_1^{N_1} (N-N_1)^{N-N_1}} \sqrt{\frac{N}{2\pi N_1 (N-N_1)}} p^{N_1} q^{N-N_1} (1 + O(\frac{1}{N})). \\ &= \left(\frac{Np}{N_1} \right)^{N_1} \left(\frac{Nq}{N-N_1} \right)^{N-N_1} \sqrt{\frac{N}{2\pi N_1 (N-N_1)}} (1 + O(\frac{1}{N})). \end{aligned}$$

$$\text{Let, } \delta = N_1 - Np \quad \Rightarrow N-N_1 = N(1-p) - \delta = Nq - \delta.$$

$$\therefore \ln \left(\frac{Np}{N_1} \right) = \ln \left(\frac{Np}{Np+\delta} \right) = -\ln \left(1 + \frac{\delta}{Np} \right)$$

$$\approx - \left(\frac{\delta}{Np} - \frac{\delta^2}{2N^2p^2} + O\left(\frac{\delta^3}{N^3p^3}\right) \right).$$

$$\& \ln \left(\frac{Nq}{N-N_1} \right) = \ln \left(\frac{Nq}{Nq-\delta} \right) = -\ln \left(1 - \frac{\delta}{Nq} \right)$$

$$\approx \left(\frac{\delta}{Nq} + \frac{\delta^2}{2N^2q^2} + O\left(\frac{\delta^3}{N^3q^3}\right) \right).$$

$$\Rightarrow \ln \left[\left(\frac{Np}{N_1} \right)^{N_1} \left(\frac{Nq}{N-N_1} \right)^{N-N_1} \right] = -(\delta + Np) \left(\frac{\delta}{Np} - \frac{\delta^2}{2N^2p^2} + O\left(\frac{\delta^3}{N^3p^3}\right) \right)$$

$$+ (Nq - \delta) \left(\frac{\delta}{Nq} + \frac{\delta^2}{2N^2q^2} + O\left(\frac{\delta^3}{N^3q^3}\right) \right)$$

$$= \delta(-1+1) + \delta^2 \left[\left(\frac{1}{2Np} + \frac{1}{2Nq} \right) - \frac{1}{Np} - \frac{1}{Nq} \right] + O(\delta^3).$$

$$= -\frac{\delta^2}{2Npq} + O(\delta^3). \quad \text{--- (II)}$$

$$\text{Also, } \sqrt{\frac{N}{2\pi N_1 (N-N_1)}} = \sqrt{\frac{N}{2\pi (\delta+Np)(Nq-\delta)}} = \sqrt{\frac{1}{2\pi Npq - 2\pi(p-q)\delta - \frac{2\pi\delta^2}{N}}}$$

$$\text{Using (I), (II) \& (III), } \Rightarrow W_N(N_1) \approx \sqrt{\frac{1}{2\pi Npq}} \exp \left(-\frac{\delta^2}{2Npq} \right) = \sqrt{\frac{1}{2\pi Npq}} \exp \left[\frac{(N_1 - Np)^2}{2Npq} \right].$$

Random Walk & Diffusion equation.

$$P_{N+1}(m) = p P_N(m-1) + q P_N(m+1). \quad \text{--- (a)}$$

$$\therefore \text{For } p = q = \frac{1}{2}, \Rightarrow P_{N+1}(m) = \frac{1}{2} P_N(m-1) + \frac{1}{2} P_N(m+1).$$

$$\text{Now, } \lim_{\tau \rightarrow 0} \frac{P_{N+1}(m) - P_N(m)}{\tau} = \frac{\partial P}{\partial t}. \quad \text{--- (b)}$$

We can also construct the difference table,

$$\begin{array}{c} P_N(m-1) \\ P_N(m) \quad P_N(m) - P_N(m-1) \\ P_N(m+1) \quad P_N(m+1) - P_N(m) \\ P_N(m+1) \end{array}$$

$$\Rightarrow \lim_{\tau \rightarrow 0} \frac{P_N(m+1) + P_N(m-1) - 2P_N(m)}{\tau^2} = \frac{\partial^2 P}{\partial x^2} \quad \text{--- (c)}$$

From (a),

$$P_{N+1}(m) - P_N(m) = \frac{1}{2} P_N(m+1) + \frac{1}{2} P_N(m-1) - P_N(m)$$

$$\Rightarrow \frac{P_{N+1}(m) - P_N(m)}{\tau} = \frac{\tau^2}{2\tau} \frac{P_N(m+1) + P_N(m-1) - 2P_N(m)}{\tau^2}. \quad \text{--- (d)}$$

\therefore Using (b) & (c) in (d),

$$\Rightarrow \frac{\partial P}{\partial t} = D \frac{\partial^2 P}{\partial x^2} \quad \text{with, } D = \frac{\tau^2}{2\tau}.$$

Poisson distribution as a special case of Binomial distribution.

$$W_N(N_1) = \frac{N!}{N_1!(N-N_1)!} p^{N_1} (1-p)^{N-N_1}.$$

$$\therefore \text{Let } W_N(N_1) = \frac{N!}{N_1!(N-N_1)!} \left(\frac{\lambda}{N}\right)^{N_1} \left(1-\frac{\lambda}{N}\right)^{N-N_1}.$$

$N \rightarrow \infty$
 $p \rightarrow 0$
 $N_1 = k$

- Large number of trials
- Probability of success in a trial is small.

Define, $\lambda = Np$.

$$P(\lambda, k) = \lim_{N \rightarrow \infty} W_N(N_1).$$

$p \rightarrow 0$
 $N_1 = k$

- Note:
- $\lim_{x \rightarrow \infty} \left(1 + \frac{1}{x}\right)^x = e$.
 - Also, $N_1 = k \ll N$.

$$= \lim_{\substack{\lambda \rightarrow 0 \\ N \rightarrow \infty}} \frac{N!}{k!(N-k)!} \left(\frac{\lambda}{N}\right)^k \left(1 - \frac{\lambda}{N}\right)^{N-k}.$$

$$= \lim_{\substack{\lambda \rightarrow 0 \\ N \rightarrow \infty}} \underbrace{\frac{N(N-1)\dots(N-k+1)}{N^k}}_{\rightarrow 1} \left(\frac{\lambda^k}{k!}\right) \left(1 - \frac{\lambda}{N}\right)^N \underbrace{\left(1 - \frac{\lambda}{N}\right)^{-k}}_{\substack{\rightarrow 1 \\ (\text{finite terms})}}.$$

$$= \frac{\lambda^k}{k!} \lim_{\frac{\lambda}{N} \rightarrow 0} \left(1 + \frac{1}{(-\frac{N}{\lambda})}\right)^{-\lambda(-\frac{N}{\lambda})}.$$

$$= \frac{\lambda^k}{k!} e^{-\lambda}.$$

$$\therefore P(\lambda, k) = \frac{\lambda^k}{k!} e^{-\lambda}$$

k trials $\ll N$.
 $\lambda = Np$, $p \rightarrow \text{small}$.